

Performance of npde for the evaluation of joint models with time to event data

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PAGE meeting – Stuart Beal methodology session

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Context

- Nonlinear Mixed Effect Models (NLMEM) increasingly more sophisticated
- Model evaluation
 - assessing the adequacy between the tested model and the data
 - important part in model development [1,2,3]
 - graphical and statistical methods available for continuous data
 - recommended methods include **visual predictive check (VPC)** and **npde** as a gold standard [4]

[1] Brendel K et al. Clin Pharmacokinet. 2007

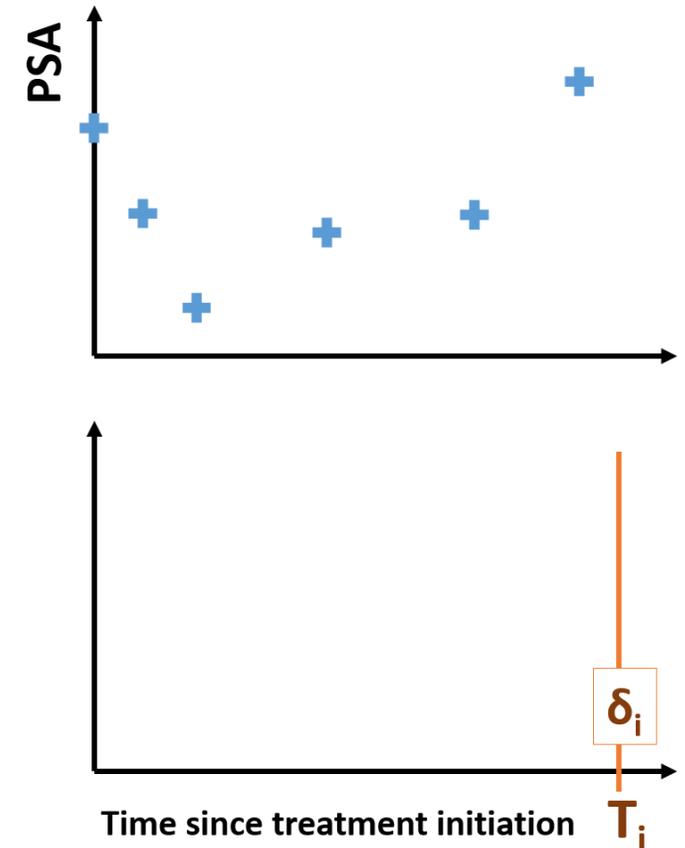
[2] FDA 1999

[3] EMA 2006

[4] Nguyen THT et al. CPT Pharmacometrics Syst Pharmacol. 2017

Joint models

- Processes of interest are followed throughout clinical trials
- Typically in oncology with biomarkers (e.g. PSA, SLD) and time-to-event (e.g. death, relapse)
- Joint models provide a promising statistical framework to estimate this association
- Support clinical decisions and treatment choices
- Increased use of joint models [1,2] with NLMEM
- How to extend npde for the evaluation of joint models ?**



[1] Sudell M et al. BMC Med Res Methodol. 2016

[2] Tardivon C et al. Clin Pharmacol Ther. 2019

Outline

- Development of npde for the evaluation of **joint model** with longitudinal and time-to-event (TTE) data
- Performance of the statistical test
- How can we visually diagnose model deficiencies ?

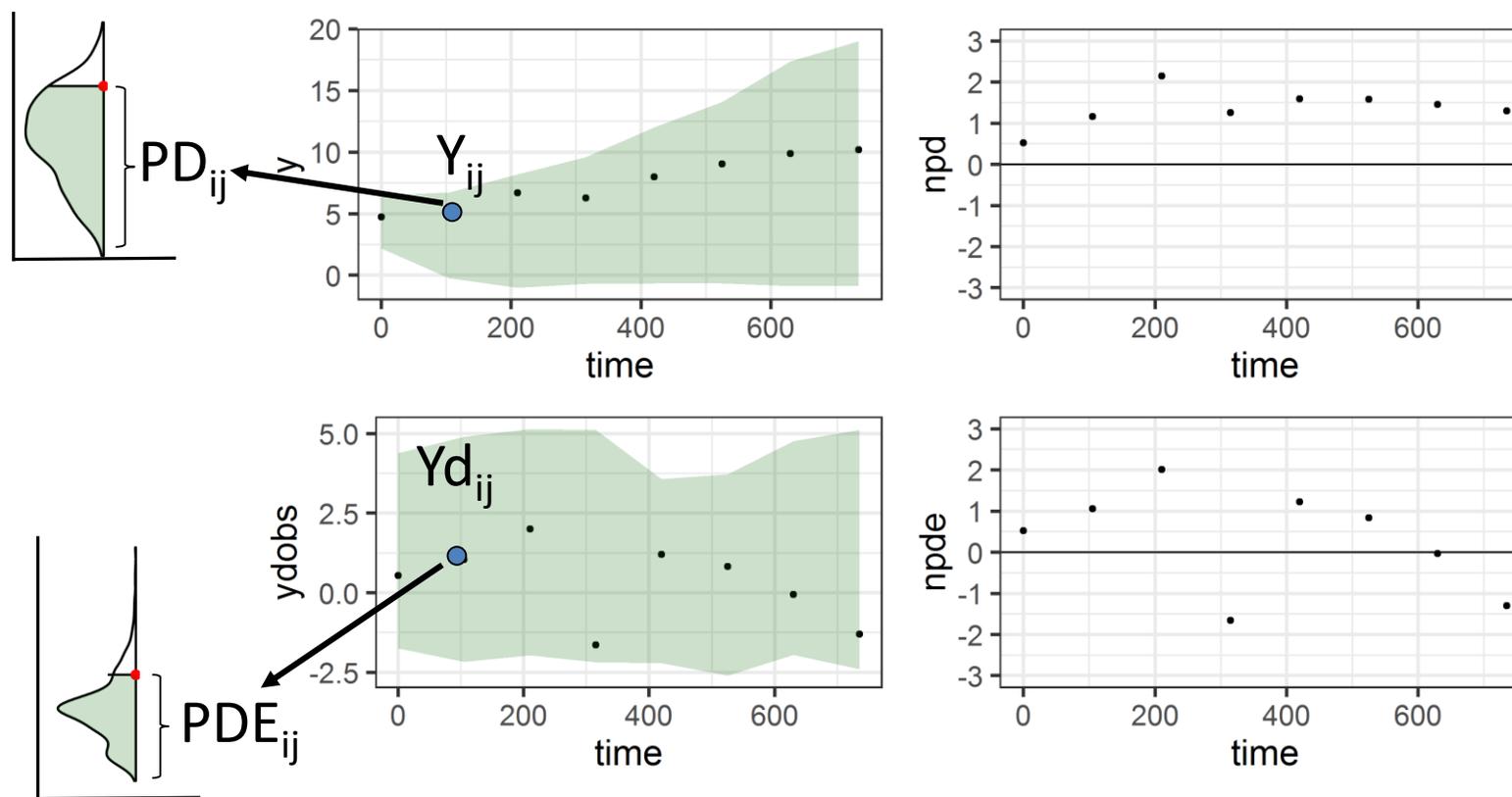
Statistical model

- y_{ij} is the j^{th} continuous observation for subject i at time t_{ij}
- T_i is the time to first event
- Model for continuous data
 - $y_{ij} = \mathbf{f}(\boldsymbol{\theta}_i, t_{ij}) + g(\theta_i, t_{ij}, \sigma)e_{ij}$, with $e_{ij} \sim \mathcal{N}(0,1)$ and g the error model
 - $\boldsymbol{\theta}_i = \mathcal{D}(\mu_L, \eta_i)$ with μ_L the fixed effects and η_i the random effects ($\eta_i \sim \mathcal{N}(0, \Omega)$)
- Dependency between observations: conditional **independence** with respect to random effects
- Model for TTE data
 - $\mathbf{h}_i(\mathbf{t}|\boldsymbol{\theta}_i) = h_0(t) \times \exp(\beta_L \cdot \mathbf{l}(\boldsymbol{\theta}_i, \mathbf{t}))$
 - with β_S the vector of parameters of the baseline hazard function h_0 , and β_L which represents the strength of the link between $\mathbf{l}(\boldsymbol{\theta}_i, \mathbf{t})$ and the hazard
 - $\mu_{TTE} = \{\beta_S, \beta_L\}$
- $\Psi = \{\mu_L, \mu_{TTE}, \Omega, \sigma\}$

Development of npde for continuous data [1]

$$pd_{ij} = F_{ij}(y_{ij}) = \int_0^{y_{ij}} p_i(y, \Psi) dy = \int_0^{y_{ij}} \int p(y|\theta_i) p(\theta_i) d\theta_i dy$$

- prediction discrepancies **pd**: quantile of an observation in its predictive distribution
- prediction distribution error **pde**: quantile of a decorrelated observation in its decorrelated predictive distribution
- normalised prediction distribution **npd**: normalization of pd
- normalised prediction distribution error **npde**: normalization of pde

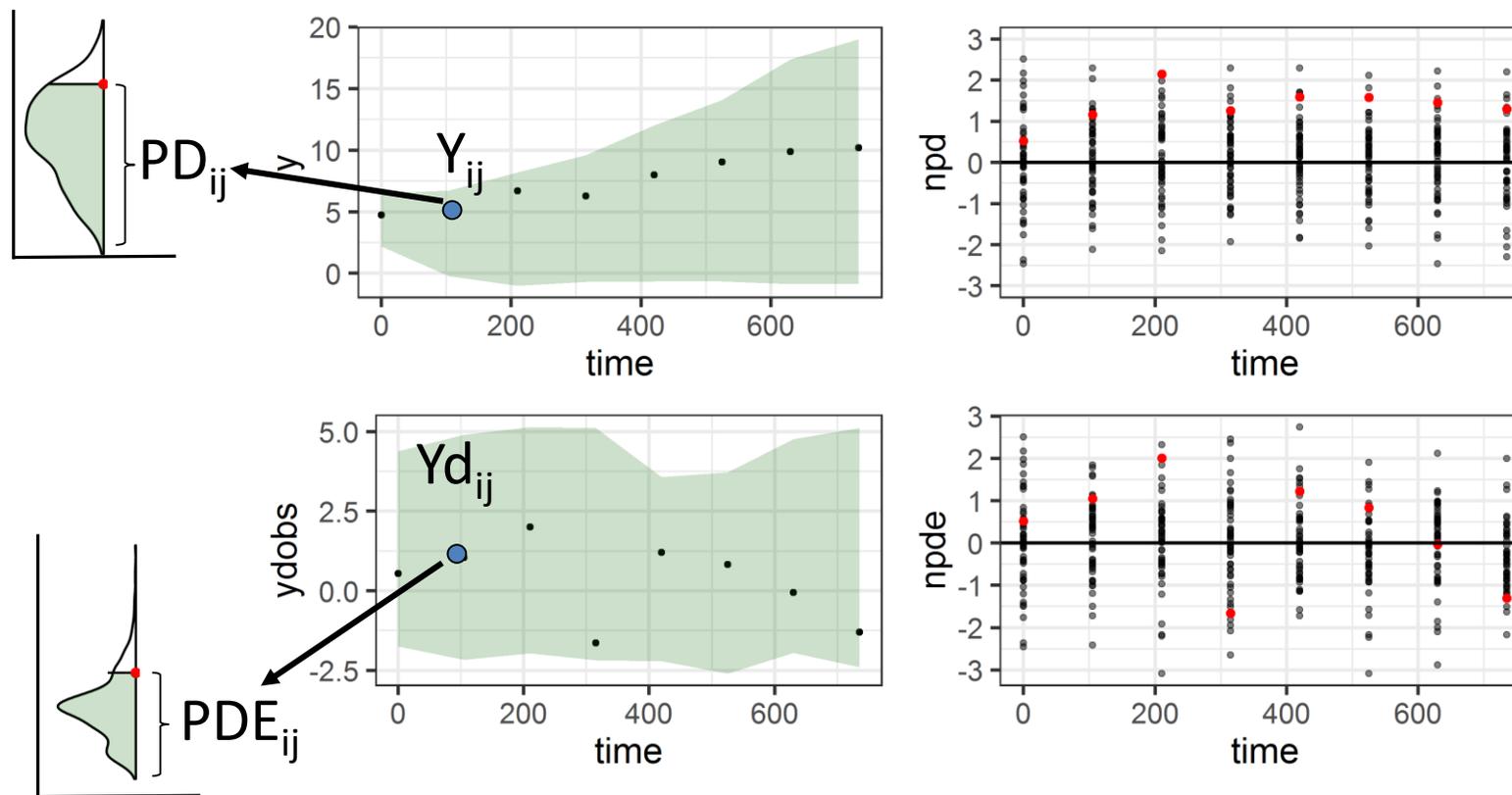


[1] Brendel K et al. Pharm Res. 2006

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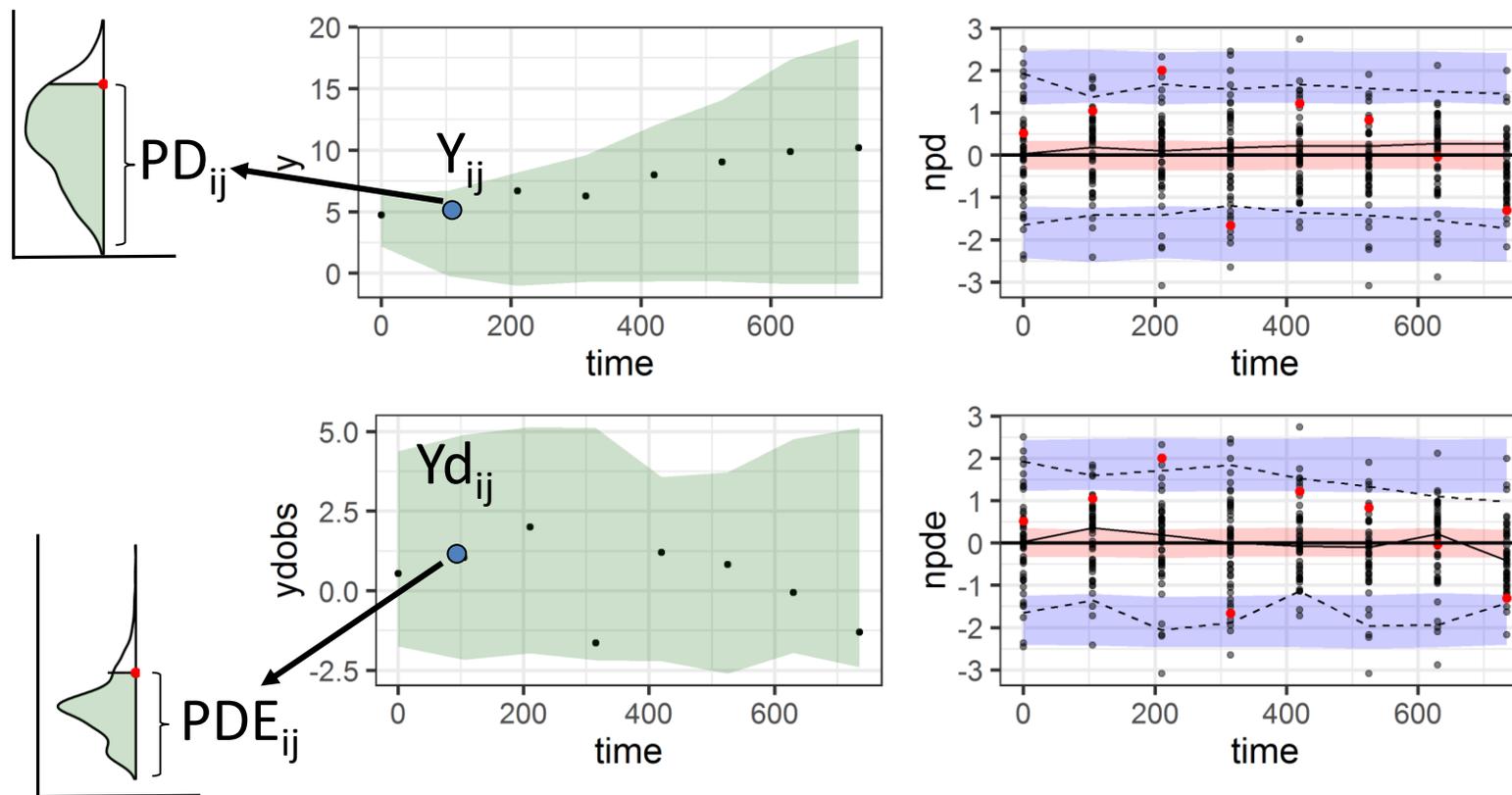


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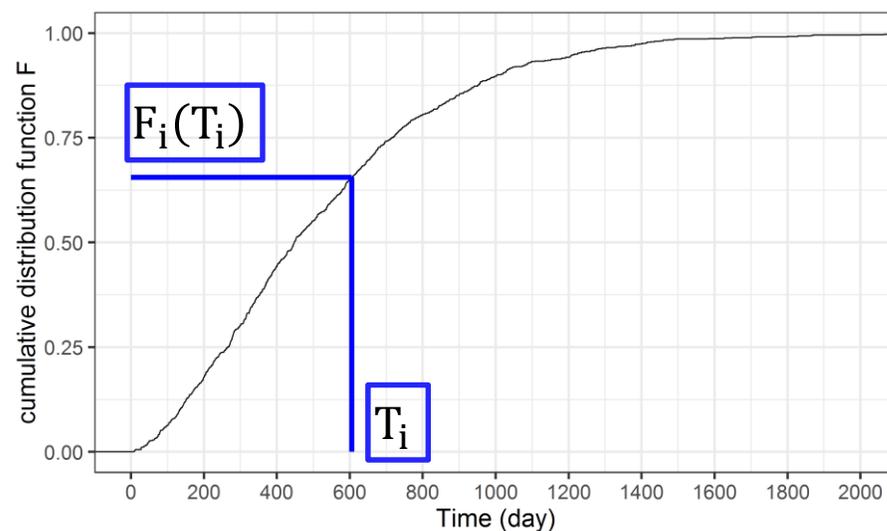


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- under H_0 :
 $pd/pde \sim U(0,1)$
 $npd/npde \sim \mathcal{N}(0,1) \rightarrow \text{test}$

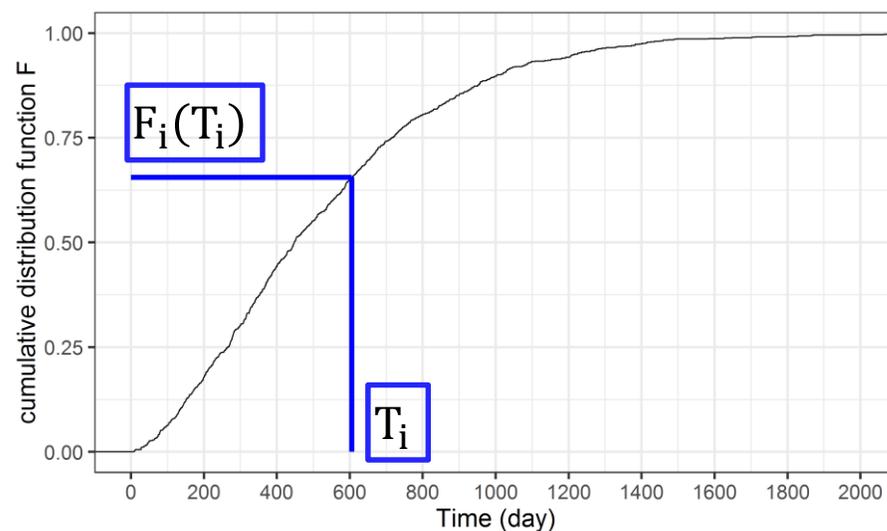
Development of npd for TTE data

- Time to event (T_i) is continuous
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 - $pd_i = F_i(T_i) = \int_0^{T_i} p_i(t, \Psi) dt = \int_0^{T_i} \int p(t|\theta_i) p(\theta_i) d\theta_i dt$



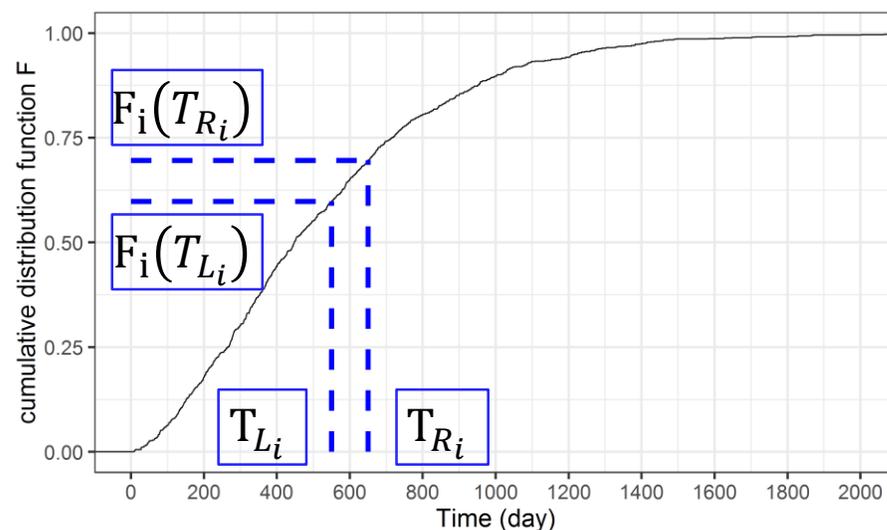
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- Time to event (T_i) can be censored (right / left / interval censored event)
 - How to deal with censoring ?



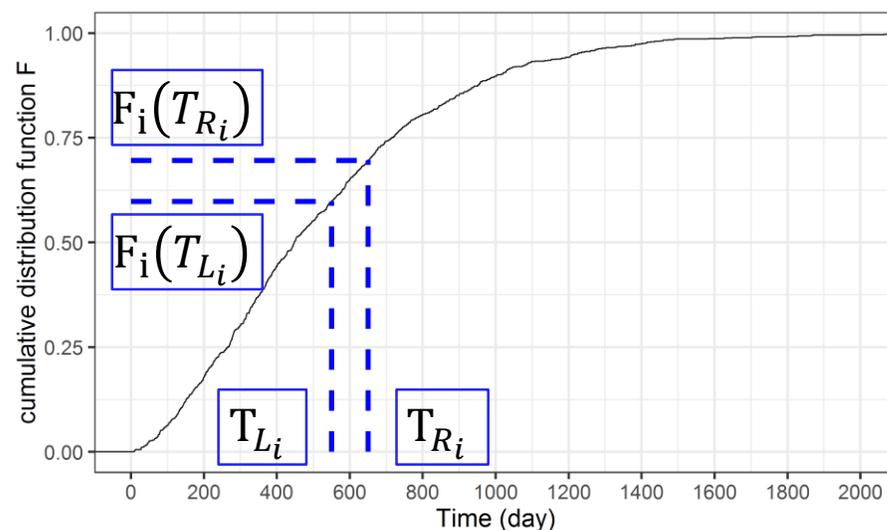
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 - How to deal with censoring ?
- pd for censored event time: same idea as [1] with npde for BLQ data:
 - T_i lies within the interval $[T_{Li}; T_{Ri}]$
 - $pd_i = U(F(T_{Li}), F(T_{Ri}))$



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In practice

- NLMEM: no analytical solution for F
- F approximated by Monte-Carlo integration

Computation of the combined test for joint model

- Computing npd / npde
 - **npde** for continuous data obtained as previously by decorrelating the observed and simulated continuous response
 - **npd** for TTE obtained using the inverse normal function of the pd (only one observation)
- Combined statistical test
 - $\Pr(\text{reject } H_0 \mid \mathbf{npde}^{(1)})$ with a Kolmogorov-Smirnov (KS) test of normality $\mathcal{N}(0,1)$
 - $\Pr(\text{reject } H_0 \mid \mathbf{npd}^{(2)})$ with a KS test of normality $\mathcal{N}(0,1)$
 - **Bonferroni correction:** $\Pr(\text{reject } H_0) = \min\left(\Pr(\text{reject } H_0 \mid \mathbf{npde}^{(1)}); \Pr(\text{reject } H_0 \mid \mathbf{npd}^{(2)})\right) \times 2$
 - rejection if $\Pr(\text{reject } H_0) < 0.05$

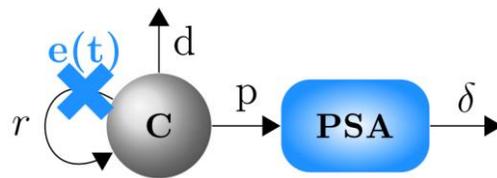
Outline

- Development of npde for the evaluation of joint model with longitudinal and TTE data
- **Performance of the statistical test**
- How can we visually diagnose model deficiencies ?

Simulation study - Model

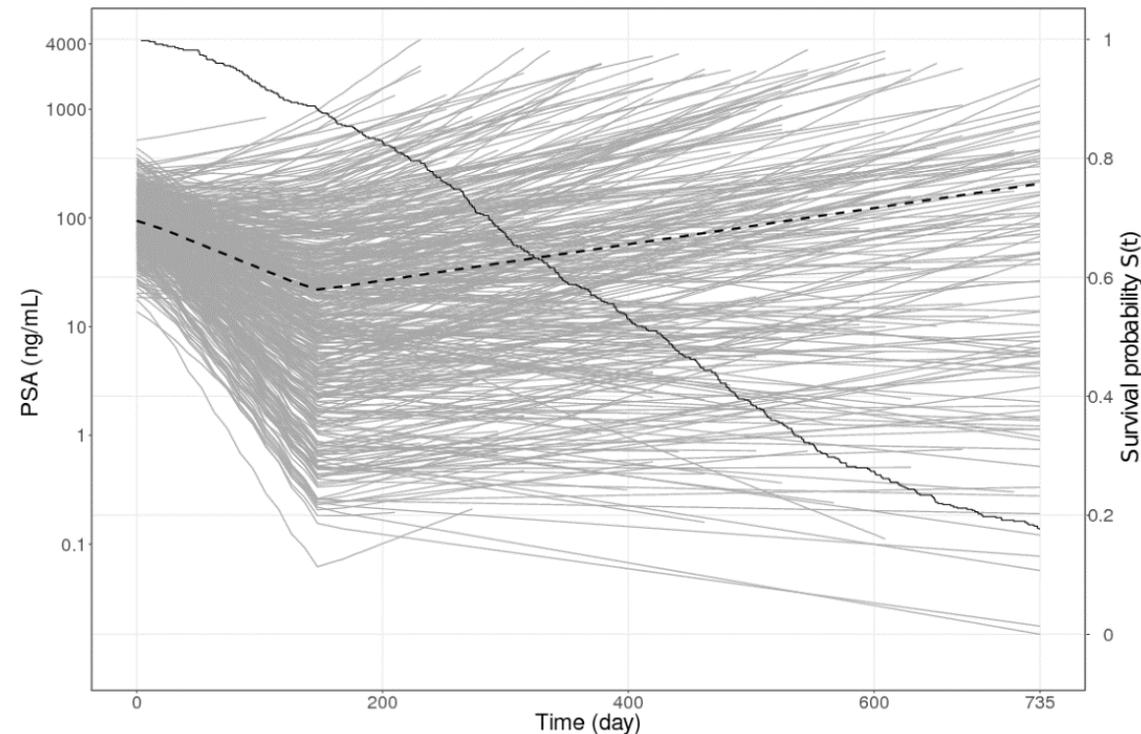
- Inspired from the work of Desmée et al. [1]
 - metastatic castration-resistant prostate cancer patient
 - original design based on a clinical trial
- Data:
 - primary outcome: **survival**
 - biomarker: Prostate Specific Antigen (PSA)

Model for PSA: biexponential



Model for the TTE outcome:

- $$h_i(t|\theta_i) = \frac{k}{\lambda} \left(\frac{t}{\lambda}\right)^{k-1} \times \exp(\beta \cdot \mathbf{PSA}(\theta_i, t))$$



Parameters	Distribution	Value	i.i.v (ω)
r (day ⁻¹)	Log-normal	0.05	0.1
PSA_0 (ng.mL ⁻¹)	Log-normal	80	0.6
ε	Logit-normal	0.3	1.5
Tesc (day)	Log-normal	140	0.6
d (day ⁻¹)	-	0.046	
δ (day ⁻¹)	-	0.23	
k	-	1.5	-
λ (day)	-	580	-
β	-	0.001	-

Simulation study

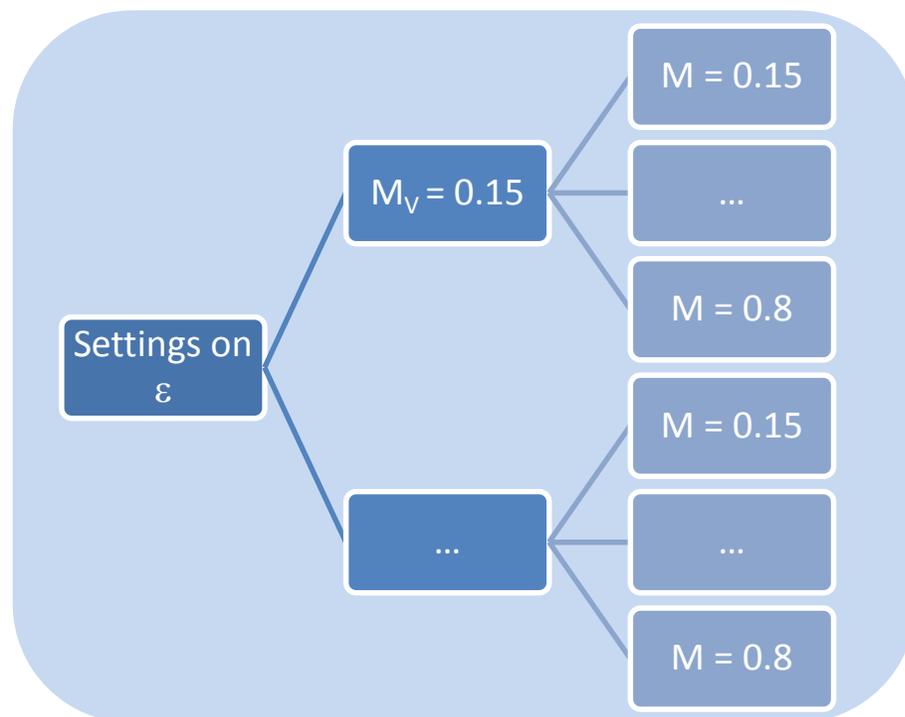
Settings

- Misspecification on **PSA model parameter**
 - $\varepsilon : \{0.15, 0.3, 0.45, 0.6, 0.8\}$
- Misspecification on **TTE model parameter**
 - $\beta : \{0, 0.001, 0.003, 0.005\}$
- Design
 - $\{50, 100, 200\}$ patients
 - follow-up censored at 735 days
 - up to 9 measurements

Simulation study

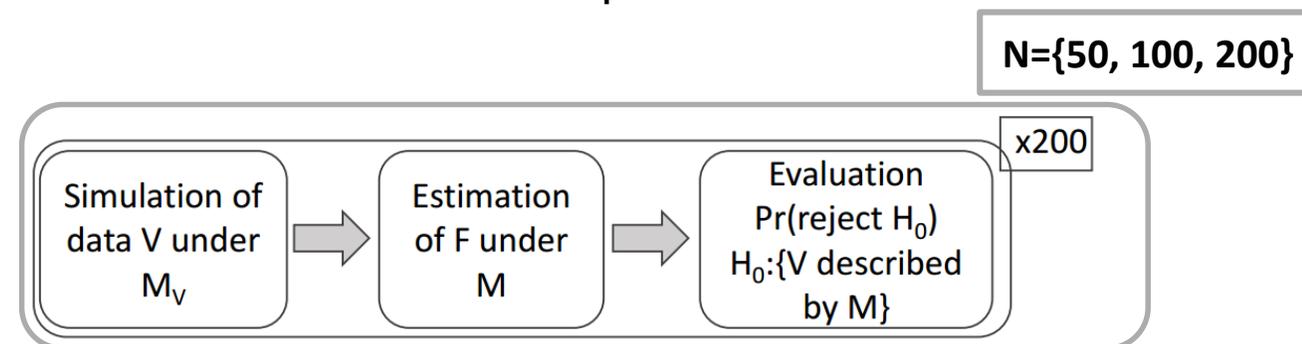
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Evaluation criteria

- Notation used:
 - V: dataset generated under model M_V (e.g. $\varepsilon = 0.15$)
 - M: model to test (e.g. $\varepsilon = 0.8$)
 - H_0 : the data V can be described by the model M
- Performance of npde:

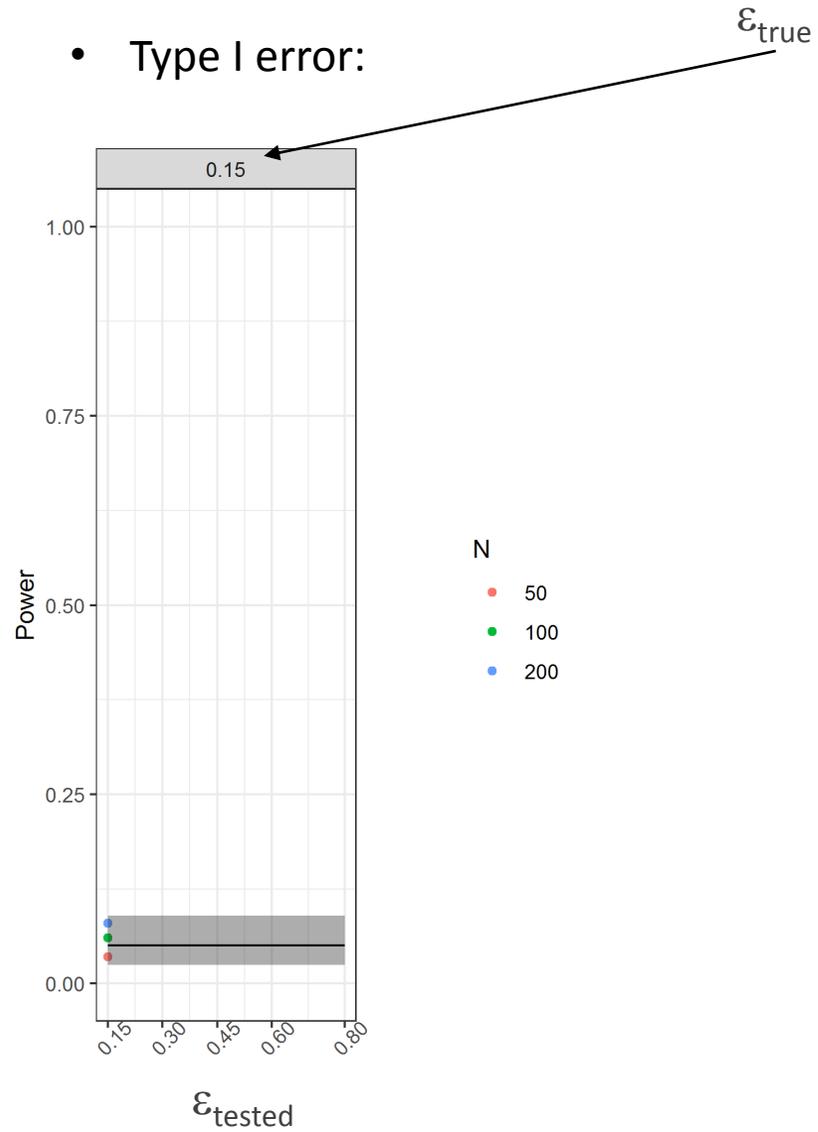


- **type I error:** % of rejection of M under H_0 ($M=M_V$)
- **power:** % of rejection of M under H_1 ($M \neq M_V$)

Simulation study – Misspecification on PSA model

Parameter: ε

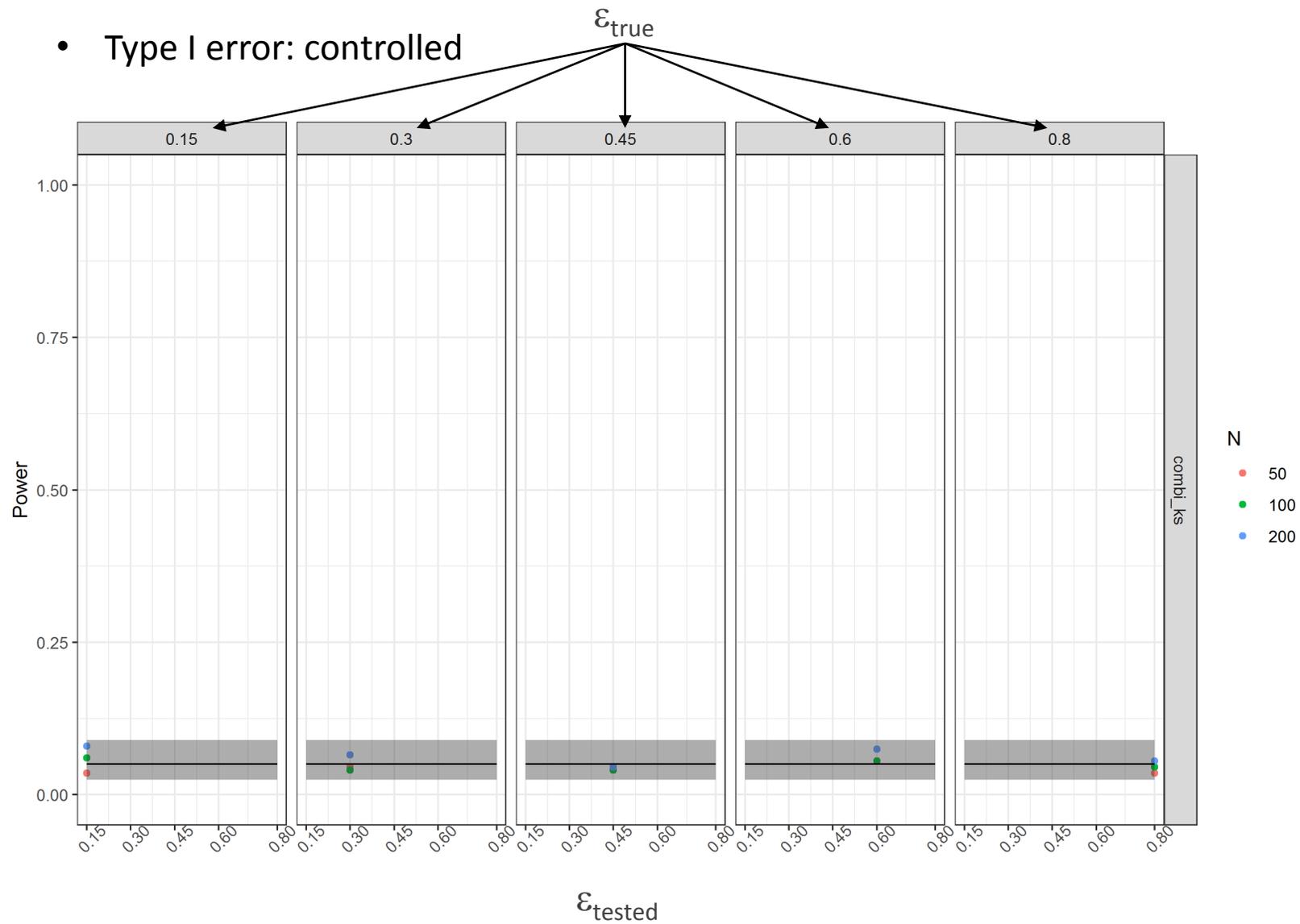
- Type I error:



Simulation study – Misspecification on PSA model

Parameter: ϵ

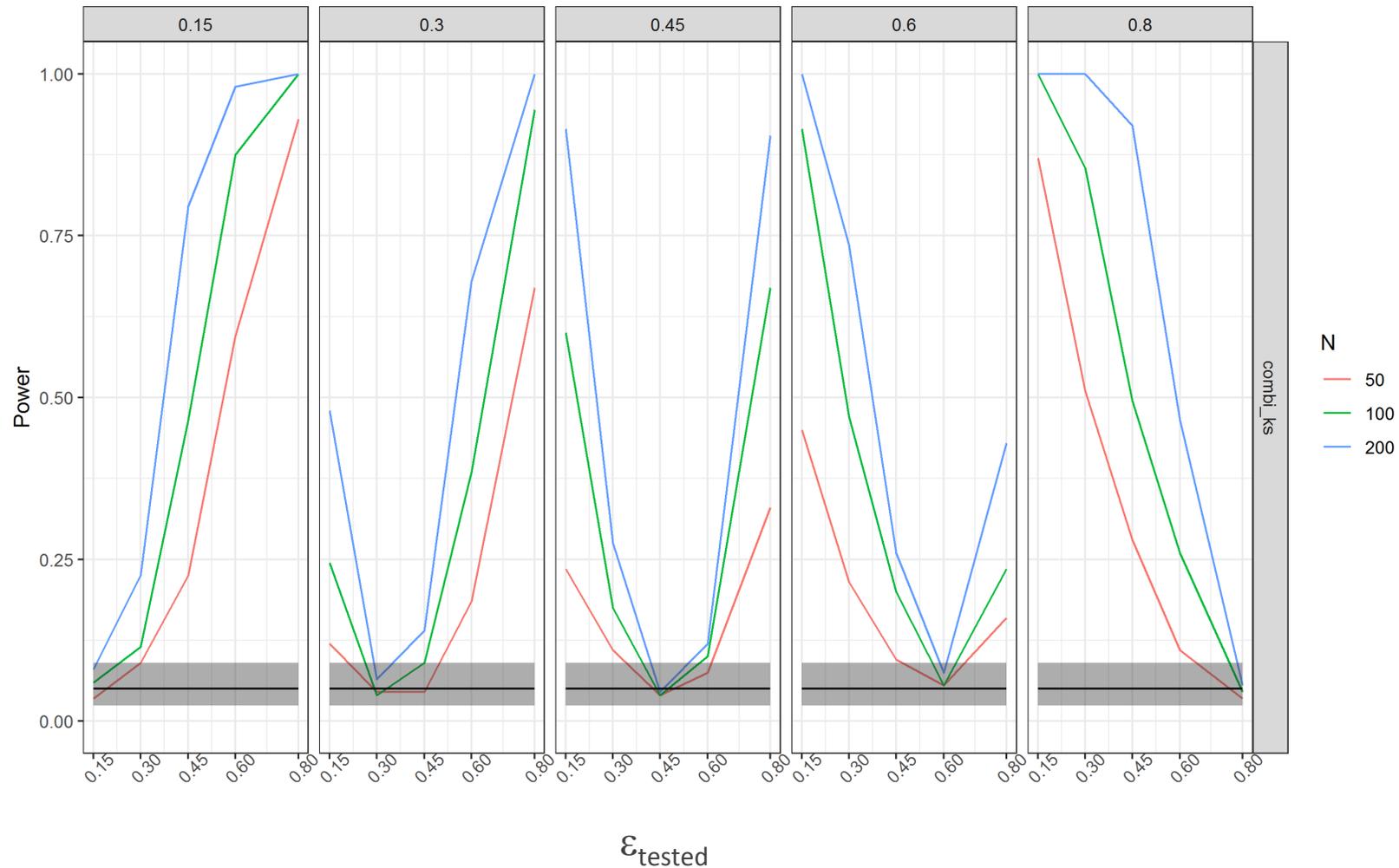
- Type I error: controlled



Simulation study – Misspecification on PSA model

Parameter: ϵ

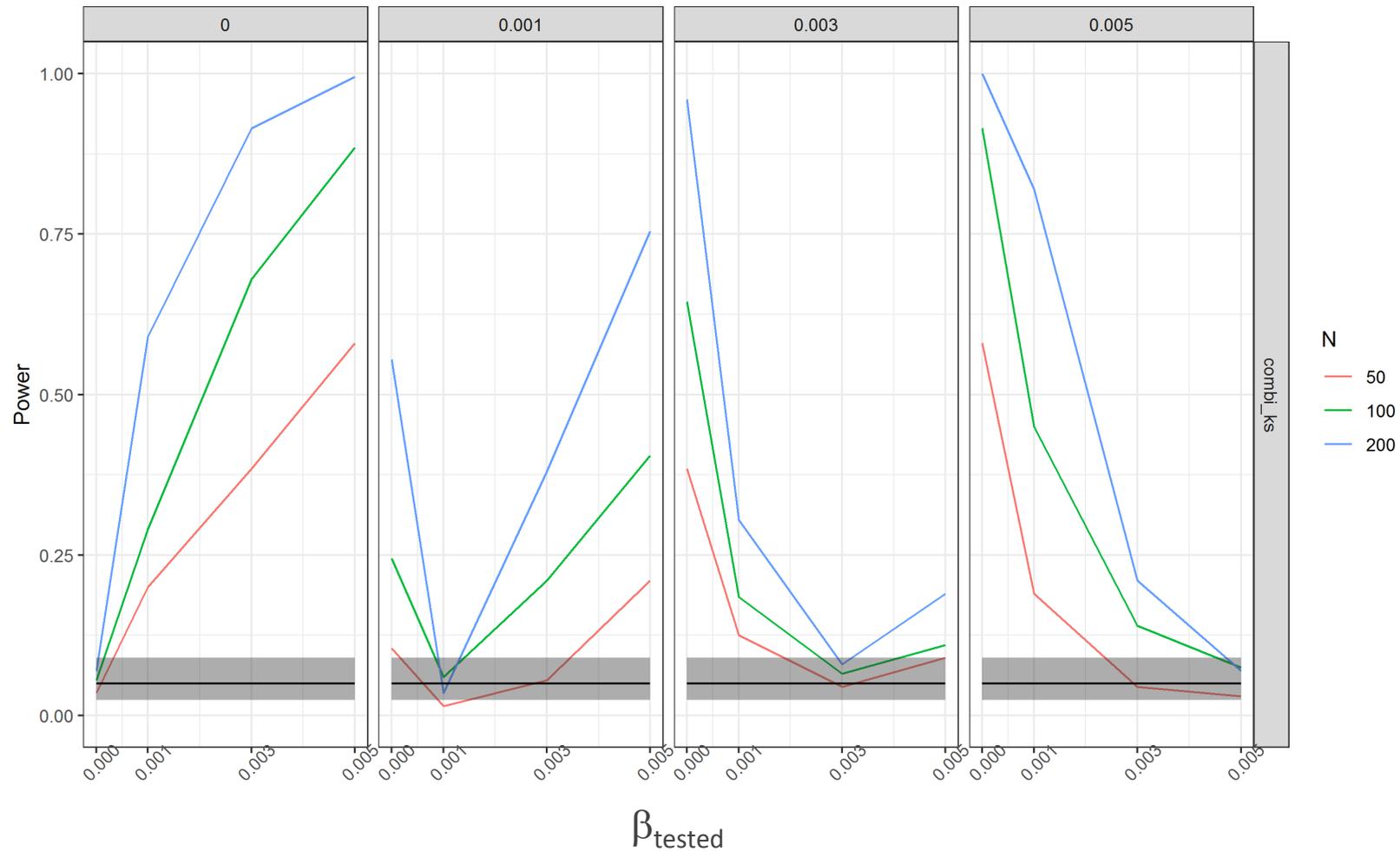
- Type I error: controlled
- Power increases as difference from true value and N increases



Simulation study – Misspecification on TTE model

Parameter: β

- Type I error: controlled
- Power increases as difference from true value and N increases



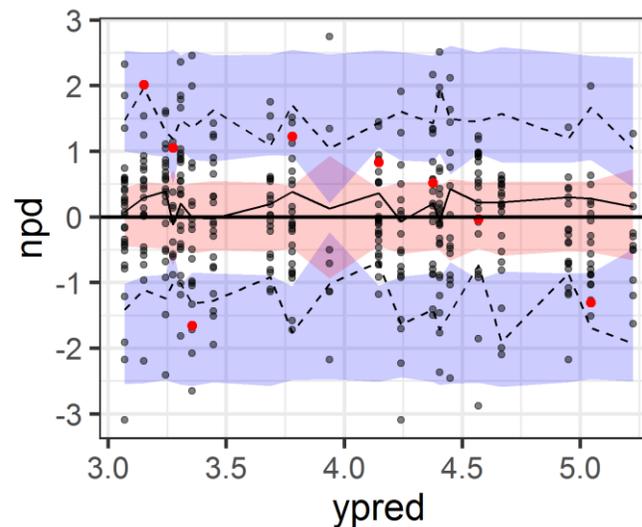
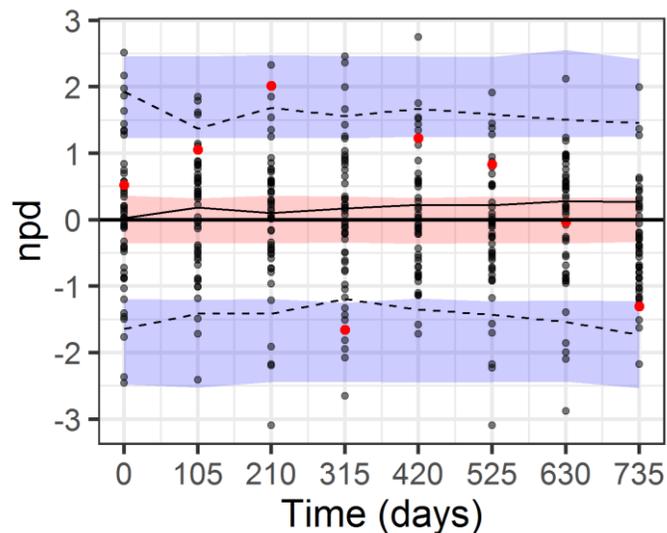
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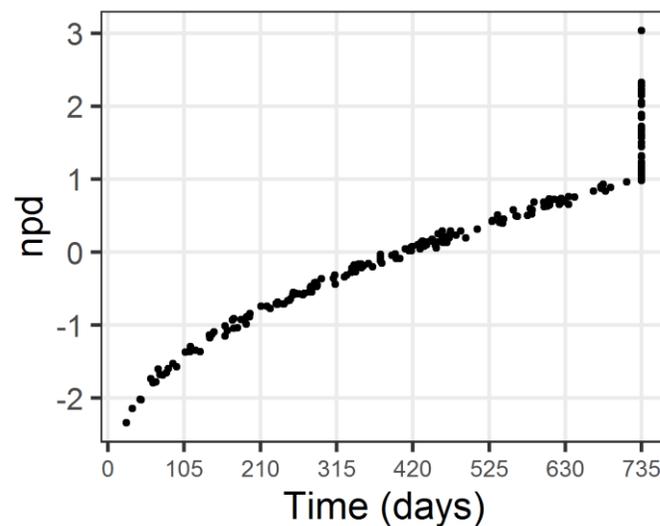
For continuous outcome:

- scatter plot of npd/npde vs time and pred

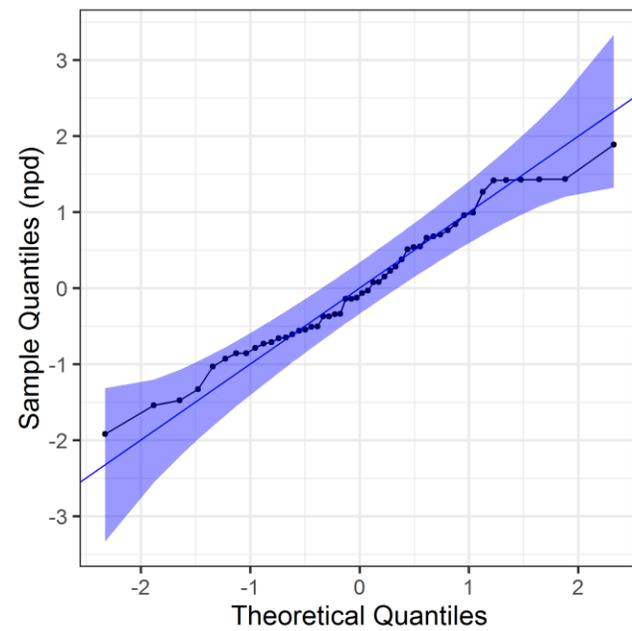


For TTE outcome:

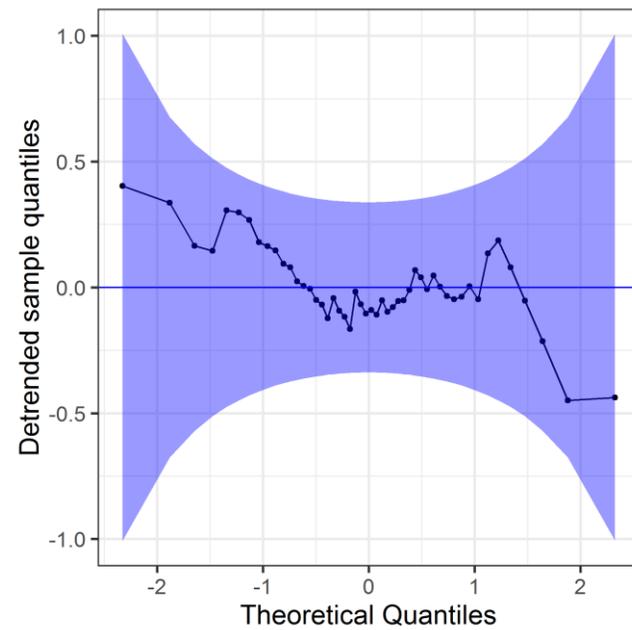
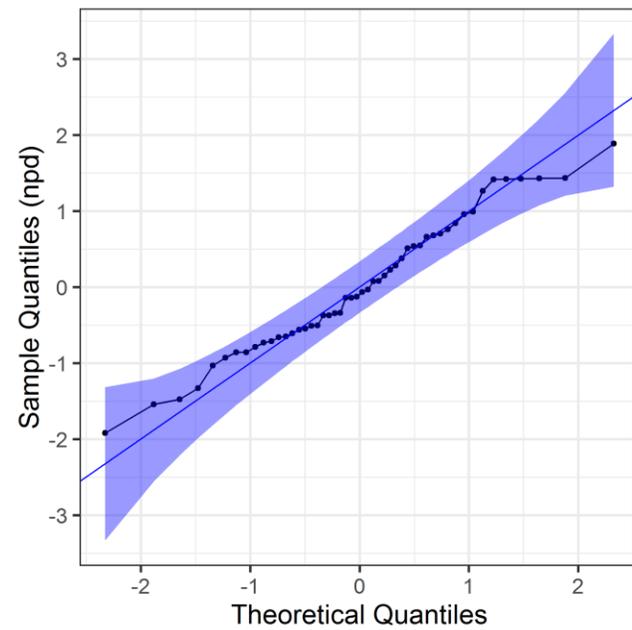
- time is the predictor itself
- if npd vs time: a trend is expected (population residuals vs. individual observations)



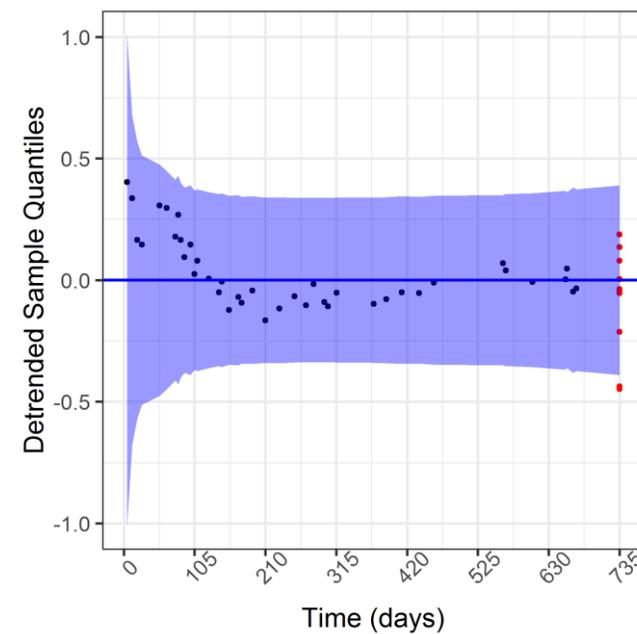
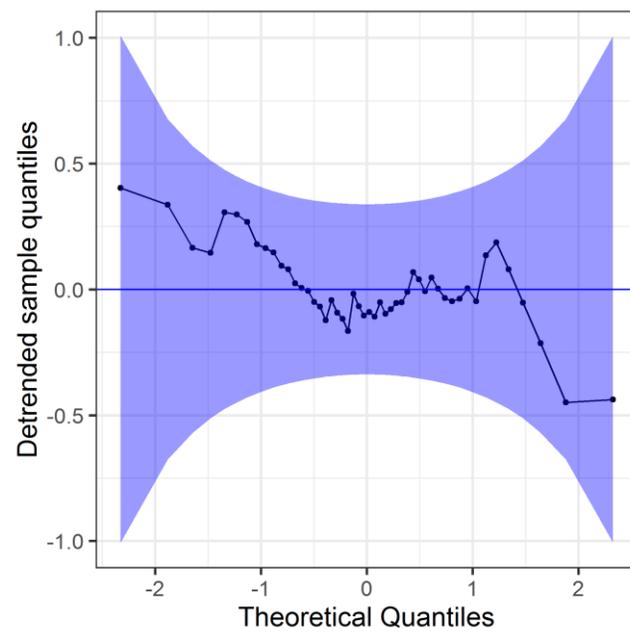
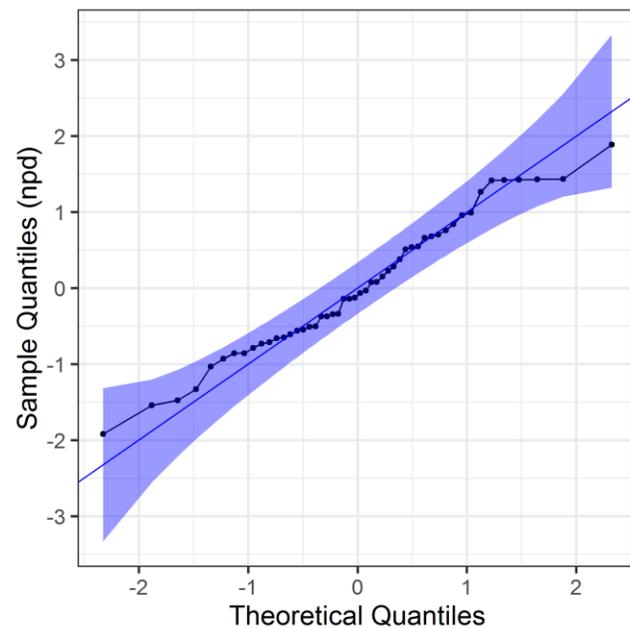
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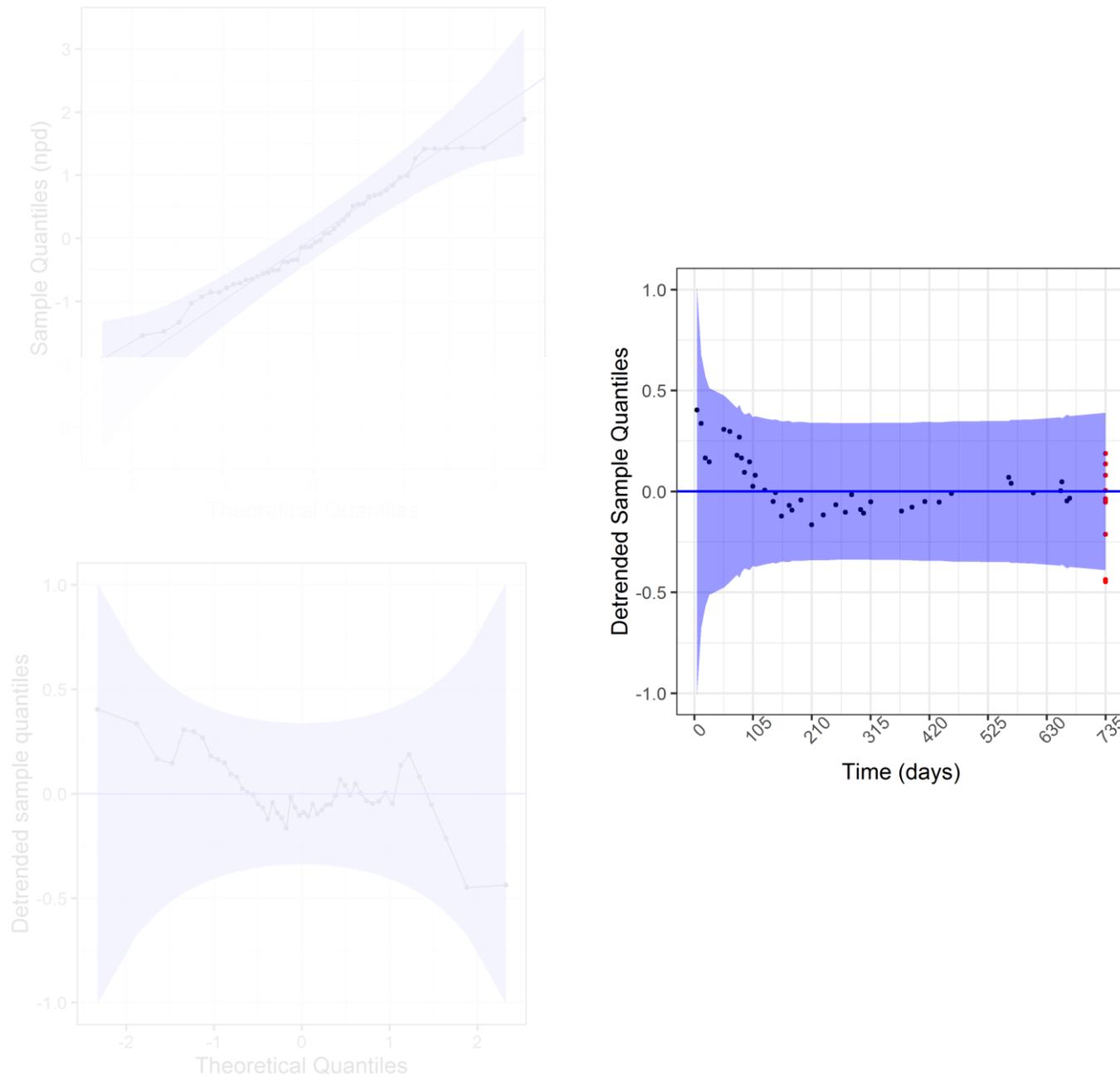
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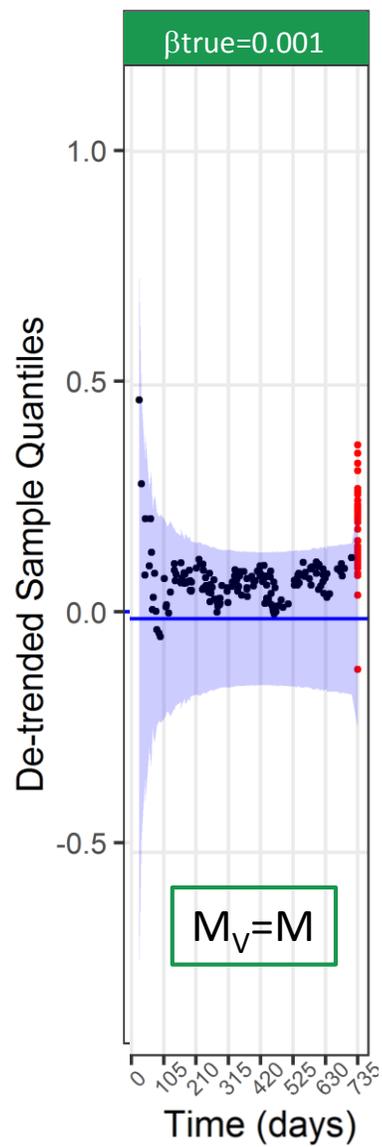


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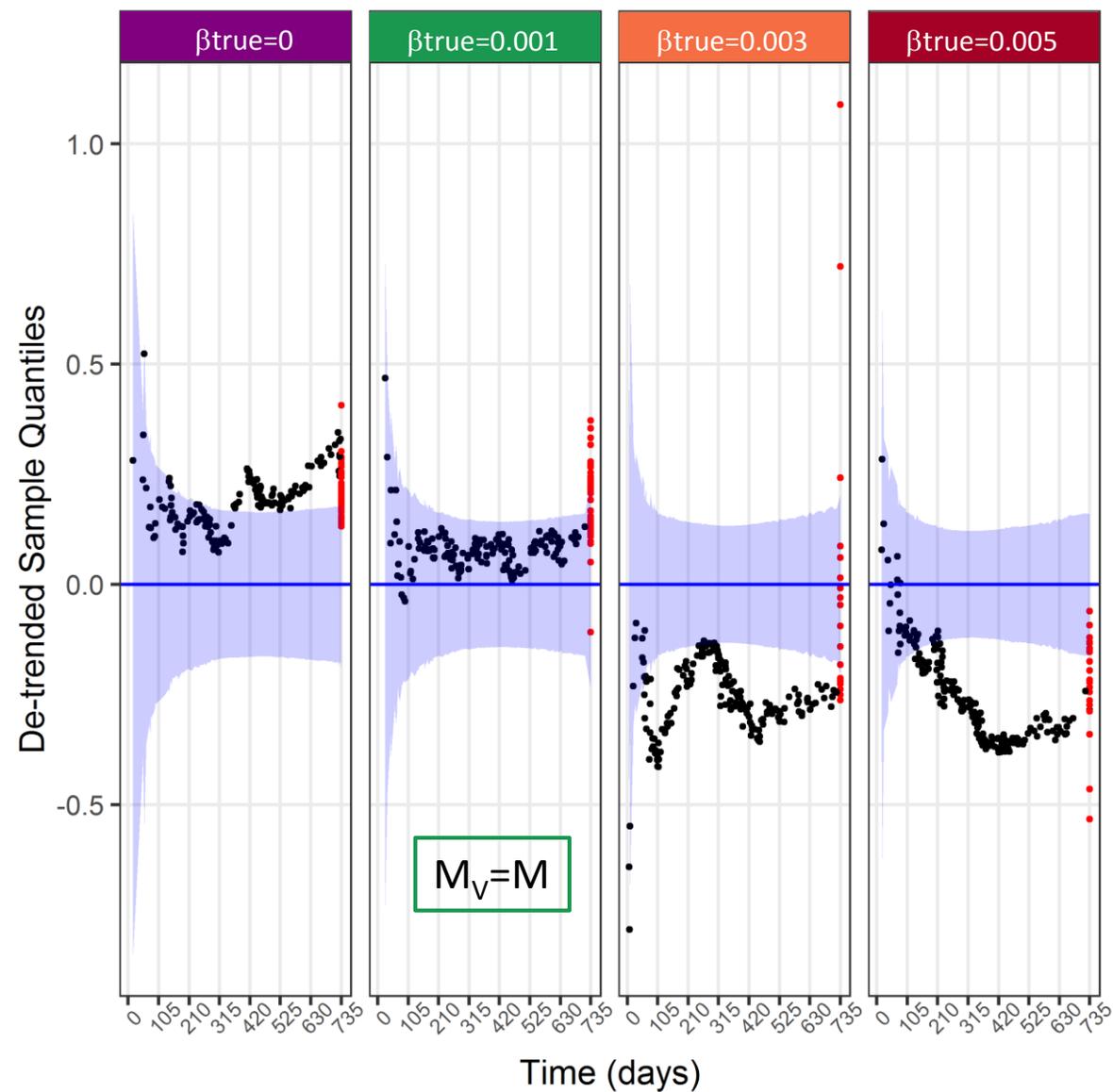
Example of graphs – TTE part

Strength of the impact of PSA on survival (β tested): 0.001



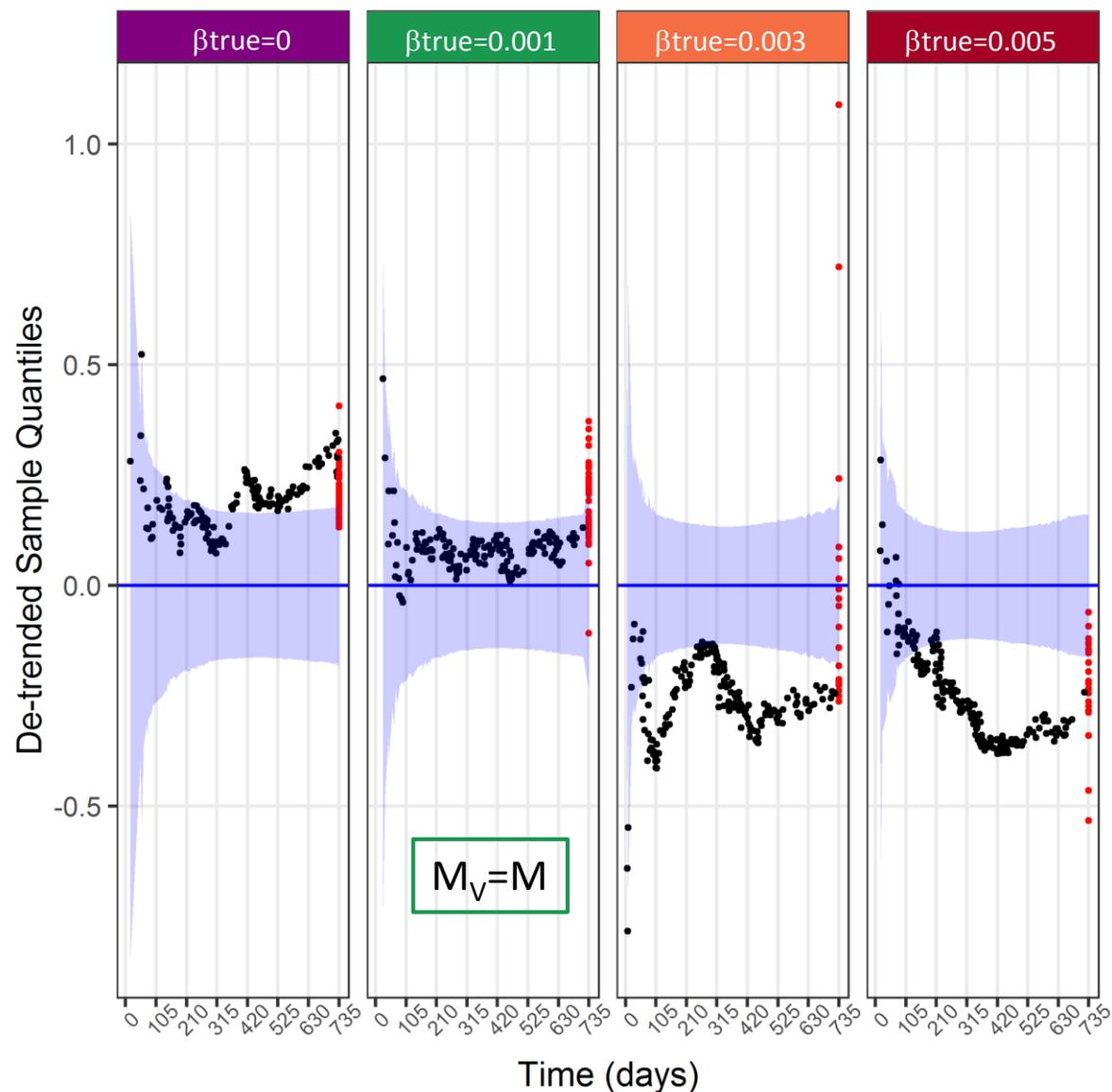
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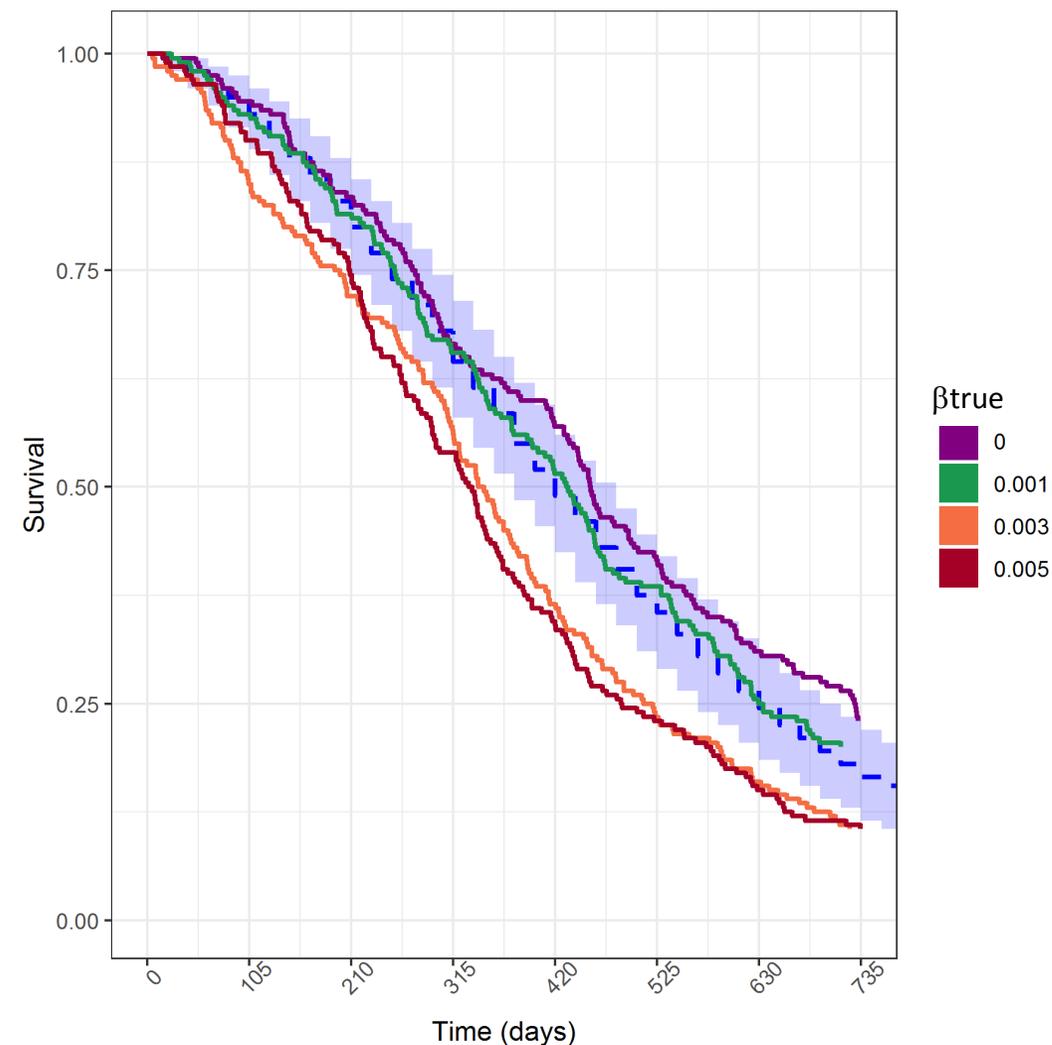


Example of graphs – TTE part

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Survival-based Visual Predictive Check



Discussion

- Good performance of the combined test
 - adequate type I error
 - power increases as the difference in shape differs
 - also evaluated in other settings
- Graphs are useful to assess the inadequacies of the model
- More details about the npd for TTE in [1]
 - good performance of npd-TTE evaluated on survival data alone, assuming longitudinal model is correct

Perspectives

- Application on data from a clinical trial
- Comparison of npd-TTE to other diagnostics for survival-type data:
 - Survival-based VPC
 - Hazard-based VPC [1]
 - Kaplan Meier Mean Covariate [2]
- Extension to repeated TTE

[1] Yeamin H and Hutmacher MM. J Pharmacokinet Pharmacodyn. 2016

[2] Hooker A and Karlsson MO. PAGE 2012

Acknowledgment

- Grant: Servier
- To my PhD supervisor
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 - Marylore Chenel
- To my Inserm colleagues
- To my Servier colleagues



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